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Association of Prehospital Time to In-Hospital Trauma Mortality in a Physician-Staffed Emergency Medicine System

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IMPORTANCE The association between total prehospital time and mortality in physician-staffed trauma systems remains uncertain.

OBJECTIVE To describe the association of total prehospital time and in-hospital mortality in prehospital, physician-staffed trauma systems in France, with the hypothesis that total prehospital time is associated with increased mortality.

DESIGN, SETTING, AND PARTICIPANTS This cohort study was conducted from January 2009 to December 2016. Data for this study were derived from 2 distinct regional trauma registries in France (1 urban and 1 rural) that both have a physician-staffed emergency medical service. Consecutive adult trauma patients admitted to either of the regional trauma referral centers during the study period were included. Data analysis took place from March 2018 to September 2018.

MAIN OUTCOMES AND MEASURES The association between death and prehospital time was assessed with a multivariable model adjusted with confounders. Total prehospital time was the primary exposure variable, recorded as the time from the arrival of the physician-led prehospital care team on scene to the arrival at the hospital. The main outcome of interest was all-cause in-hospital mortality.

RESULTS A total of 10 216 patients were included (mean [SD] age, 41 [18] years; 7937 men [78.3%]) affected by predominantly nonpenetrating injuries (9265 [91.5%]), with a mean (SD) Injury Severity Score of 17 (14) points. Of the patients, 6737 (66.5%) had at least 1 body region with an Abbreviated Injury Scale score of 3 or more. A total of 1259 patients (12.4%) presented in shock (with systolic pressure <90 mm Hg) and 2724 (26.9%) with severe head injury (Abbreviated Injury Scale score \geq 3 points). On unadjusted analysis, increasing prehospital times (in 30-minute categories) were associated with a markedly and constant increase in the risk of in-hospital death. The odds of death increased by 9% for each 10-minute increase in prehospital time (odds ratio, 1.09 [95% CI, 1.07-1.11]) and after adjustment by 4% (odds ratio, 1.04 [95% CI, 1.01-1.07]).

CONCLUSIONS AND RELEVANCE In this study, an increase in total prehospital time was associated with increasing in-hospital all-cause mortality in trauma patients at a physician-staffed emergency medical system, after adjustment for case complexity. Prehospital time is a management objective in analogy to physiological targets. These findings plead for a further streamlining of prehospital trauma care and the need to define the optimal intervention-to-time ratio.

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Supplemental content

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Group Information: The French Trauma Research Initiative members appear at the end of the article.

Corresponding Author: Tobias Gauss, MD, Department of Anesthesia and Critical Care, Hôpital Beaujon, Hôpitaux Universitaires Paris Nord Val de Seine, Assistance Publique-Hôpitaux de Paris, 100 Boulevard du Général Leclerc, 92110 Clichy, France (tgauss@protonmail.com). he concept of time as essential determinant of outcome in the early care of trauma patients is a prevailing doctrine,^{1,2} captured in the term *the golden hour* of trauma.^{3,4} The concept of trauma as a time-sensitive condition has guided education and policy for decades in the international trauma community. Despite this cultural penetrance,⁵ evidence in favor remains elusive, and the concept is controversial.⁶⁻¹⁰

The concept of trauma as a time-sensitive health condition was developed and nurtured in a specific emergency medical system (EMS), mainly a paramedic-based system in North America. The international trauma community transposed this concept to diverse, specific settings, even if the respective context was different from that of the North American system.¹¹ Data on system performance from physician-staffed EMS are far less abundant than data from EMS systems. On the one hand, physician-staffed EMS are thought to add complexity to a complex situation by performing more interventions that result in longer total prehospital time (TPT).^{12,13} On the other side, proponents of the physician-led critical care team advocate a more tailored provision of advanced prehospital critical care. The ensuing debate oscillates between dichotomous concepts such as scoop and run or stay and play.^{11,13,14} The reality is probably less easy to frame.

Based on this rationale and considering the comparative lack of data for physician-staffed EMS, it appeared justified to reconsider the association of prehospital total time with inhospital mortality in a physician-staffed EMS. The objective was to describe the association of total prehospital time and in-hospital mortality in a large and representative cohort of trauma patients from 2 regional and very geographically distinct trauma systems in France (1 urban and 1 rural and mountainous). We hypothesized that a longer TPT is associated with a mortality increase in a physician-staffed EMS system.

Methods

This study follows the Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) guidelines. A checklist for this cohort study is provided in eTable 1 in the Supplement.

Ethical Approval

This is an observational study using data from 2 prospective multicenter regional trauma registries, the TraumaBase, from the Paris area, Île-de-France, and Trauma System of the Northern French Alps Emergency Network (*Trauma System du Réseau Nord Alpin des Urgences* [TRENAU]), from the Rhône-Alpes region. Both registries have obtained approval from the institutional review board (*Comité de Protection des Personnes*, Paris VI and Clermont-Ferrand), the Advisory Committee for Information Processing in Health Research (*Comite Consultatif Pour le Traitement de l'Information en Matière de Recherche Dans le Domaine de la Santé*, 11.305bis and 15.038bis), and from the National Data Protection Agency (*Commission Nationale de l'Informatique et des Libertés* 911461 and 915372), waiving the need for informed consent. Both registries use al-

Key Points

Question How are prehospital time and mortality associated in a physician-staffed trauma system?

Findings The results of this cohort study from 2 French trauma registries demonstrate a linear association between total prehospital time and in-hospital all-cause mortality. The odds of death increased by 8% for each 10-minute increase in prehospital time.

Meaning These findings call for a further streamlining of prehospital trauma care and the need to define the optimal intervention-to-time ratio.

gorithms for consistency, coherence, and professional data monitoring. Data monitoring for the TraumaBase is assured by the Biostatistics Laboratory of Paris 7 and for the TRENAU registry by professional statisticians employed by the network.

Setting

The structure and triage in both trauma systems have been previously described.^{15,16} In summary, physician-staffed mobile intensive care units provide the prehospital care in both systems. A mobile intensive care unit team provides the equivalent of advanced trauma life support-based prehospital interventions (eg, sedation, prehospital anesthesia, advanced airway management, chest decompression). The TraumaBase in Île-de France covers an urban area of 4633.2 mi² (12 000 km²) with a population of 12 million inhabitants and 15 million visitors per year. Almost exclusively, 8 physician-staffed EMS (Service Aide Medicale Urgente [SAMUs]) organized into 43 mobile intensive care units, 37 SAMUs, 6 Paris Fire Brigades (Brigade de Sapeurs de Pompiers), and 6 designated level I centers provide the trauma care; no lower-level care centers have been designated. The TRENAU area covers an area of 6949.8 mi² (18 000 km²) in the Northern French Alps, with a population of 2 million inhabitants and a high seasonal variation (8 million tourists each year). Care in the TRENAU network is provided by 3 SAMUs and a total of 14 designated trauma hospitals, with 2 level I, 1 level II, and 11 level III centers. The TRENAU network uses specific management and triage guidelines.¹⁵ Within the Paris network, national triage¹⁷ and management guidelines apply, such as those for shock.¹⁸ At any given time, clinical management was left to the discretion of the responsible physician (prehospital or in-hospital). Both networks and registries are financially supported by the respective Regional Health Authorities (Agence Régionale de Santé) in Île-de-France and Rhône-Alpes.

Study Populations

All consecutive trauma patients triaged to 1 of the 6 regional, designated Paris level I trauma centers or 1 of the 14 centers participating in the TRENAU network were screened for inclusion. Patients who were not managed by a physician-staffed mobile unit or transported to a participating hospital (ie, recorded in the data set as *death on scene* or *not transported*) or transferred from a hospital outside of the respective system were not included. Only patients directly admit-

ted to level I or level II centers (if in TRENAU) were included. The enhanced care teams in both systems are able to provide a full spectrum of advanced trauma life support care, including airway management with rapid sequence induction for intubation, sedation and analgesia, fluid therapy, vasoactive medication, fracture and dislocation reduction, chest decompression, hemorrhage control (eg, by tourniquets, splinting, compression), red blood cell concentrate transfusion, and other techniques. They currently do not provide thoracotomy or blood-component therapy other than red blood cell concentrates, which occur in less than 3% of cases in both systems.

We defined early critical prehospital trauma interventions as any of the following: prehospital intubation, vasopressor use, osmotherapy, and fluid resuscitation of more than 1000 mL. Shock was defined as a systolic arterial pressure less than 90 mm Hg.

eTable 1 in the Supplement provides the complete list of epidemiological, clinical (prehospital and in-hospital), physiological, and biological variables that were recorded for each patient. In brief, both data collections followed the revised version of the Utstein Template for Uniform Reporting of Data Following Major Trauma.¹⁹ Calculation of the 2005 version of Injury Severity Score (ISS) completed this information.

Exposure Variable and Outcome Criteria

Total prehospital time was the primary exposure variable, recorded as the time of arrival of the physician-led prehospital care-team on scene to the time of arrival at the hospital. Arrival on scene, leaving the scene, and arrival at the admitting hospital are systematically recorded in both registries based on prehospital intervention run sheets and/or dispatch center data; the exact timing of injury was not recorded in either registry.

The main outcome of interest was all-cause in-hospital mortality. Patients who died within 24 hours were retained in the analysis to minimize survivor bias. Cause-specific death, including death from bleeding and death from head injury, was considered the secondary end point. In both registries, the causes of death were based on a judgment by a physician reviewing all evidence at routine case closure. The physician performing case closure was not part of the working committee of this study.

Statistical Analysis

Sample Size

The retrospective nature of the study predetermines the sample size. A post hoc power calculation was performed using a 1-sample correlation Fisher z test. We used the regression coefficient of total prehospital time as an alternative hypothesis and a coefficient of 0 corresponding to the null hypothesis with no association with traumatic death.

Statistical Methods

Continuous variables were presented with means and SDs or with median and interquartile range when appropriate. Categorical variables were presented with frequencies and percentages.

The association between death and prehospital time was assessed graphically by plotting the risk of death on prehospital time. A generalized linear mixed model (GLMM) with random effect by registry and EMS system explored the association between death and prehospital time with a logit function. We assessed the departure of linearity in the association between death and time graphically and added the polynomial terms in the GLMM if needed (eTable 6 in the Supplement). Homogeneity between categories was tested using the χ^2 test.

We fitted the best model according to likelihood ratio test and Akaike information criterion. In addition, a generalized additive model (GAM) assessed and explored the linearity of the association between death and time. The GAM is a flexible model of GLMM with a link function based on smoothing splines. The amount of smoothing is controlled by the number of equivalent degrees of freedom. We graphically assessed the smooth function to fix the equivalent degrees of freedom for prehospital time. We assessed the deviance variation that occurred depending on the use of the linear term or the smoothing term in the equation.

To control for potential variables associated with case complexity, a GLMM and GAM model was applied in multivariable analysis to adjust for the confounders age, sex, severity of anatomic injury pattern (by ISS), systolic blood pressure, and Glasgow Coma Scale (GCS) score in a parsimonious method. These 5 variables are strongly correlated with prehospital interventions, such as airway management. The more severely a patient is in shock (by systolic blood pressure measurement) or the greater the likelihood of a traumatic brain injury (TBI) (as measured by the GCS), the more severely injured the patient is (per the ISS). The high injury load may trigger more interventions and longer prehospital times. In consequence, for the sake of simplicity and to avoid collinearity, the models did deliberately not include prehospital interventions.

Two-sided *P* values <.05 were considered significant. All statistical analyses were performed using Stata SE version 14.0 (StataCorp).

Missing Data

Multiple imputation of missing values to assess total prehospital time was not feasible. Several variables were not routinely collected in both registries (eg, distance between scene of injury and hospital, dispatching decision). Measurement errors in variables that are not routinely collected could lead to inappropriate assessments when applied to multiple imputation. Data were unlikely to be missing at random, and this could lead to a bias toward the null when applying multiple imputation. For this reason, a complete case analysis was performed. Comparability of patients with and without missing data on TPT were assessed in terms of demographic characteristics, prehospital condition, injury severity, and inhospital death.

Results

In total, 6441 patients were screened for inclusion from January 2011 to December 2016 in the TraumaBase database and 8888 were screened for inclusion from the TRENAU database from January 2009 to December 2016. The main reason to exclude pa-

	No. (%)				
Characteristic	Total	Paris, Île-de-France (TraumaBase)	Northern French Alps (TRENAU)		
No.	10126	5067	5059		
Age, mean (SD), y	41 (18)	38 (17)	42 (18)		
Male sex	7937 (78.3)	3998 (78.9)	3939 (77.9)		
Penetrating injury	861 (8.5)	568 (11.2)	293 (5.8)		
Circumstances					
Road traffic injury	5598 (55.2)	3014 (59.5)	2549 (50.4)		
Fall	3053 (30.1)	1231 (24.3)	1805 (35.7)		
Stabbing	605 (5.9)	399 (7.9)	201 (4.0)		
Gun shot	250 (2.5)	169 (3.3)	80 (1.6)		
Prehospital systolic blood pressure, mm Hg					
Mean (SD)	117 (32)	109 (34)	126 (2.5)		
<90 mm Hg	1259 (12.4)	946 (18.7)	313 (6.2)		
Prehospital Glasgow Coma Scale score					
3-8	1518 (15.0)	889 (17.5)	629 (12.4)		
9-13	963 (9.5)	517 (10.2)	446 (8.8)		
13-15	7453 (73.6)	3648 (72.0)	3805 (75.2)		
Injury Severity Score					
Mean (SD)	17 (14)	18 (15)	17 (13)		
0-15	5313 (52.5)	2679 (52.9)	2634 (52.1)		
16-24	2034 (20.1)	1013 (20.0)	1021 (20.2)		
25-34	1757 (17.4)	818 (16.1)	939 (18.6)		
≥35	966 (9.5)	554 (10.9)	412 (8.1)		
Overall Abbreviated Injury Scale score ≥3	6737 (66.5)	3302 (65.2)	3435 (67.9)		
Severe head injury (Abbreviated Injury Scale-head score ≥3)	2724 (26.9)	1346 (26.6)	1378 (27.2)		
Secondary transfer					
In-hospital mortality	968 (9.6)	566 (11.2)	402 (7.9)		

Abbreviation: TRENAU, *Trauma System du Réseau Nord Alpin des Urgences*.

tients was missing data (eFigure 1 and eTable 2 in the Supplement). In total, 10126 patients (5059 from TRENAU and 5067 from TraumaBase) were retained for complete case analysis.

Patients had a mean (SD) age of 41 (18) years, and most were men (7937 [78.3%]) affected by predominantly nonpenetrating injury (9265 [91.5%]), with a mean (SD) ISS of 17 (14) points. In 6737 patients (66.5%), at least 1 body region had an AIS score of 3 points or more. A total of 1259 patients (12.4%) presented in shock (systolic pressure <90 mm Hg), and 2724 (26.9%) presented with severe head injury (AIS \geq 3 points; **Table 1**).

The median TPT was 73 (interquartile range [IQR], 54-100) minutes in the TraumaBase cohort and 60 (IQR, 45-80) minutes in the TRENAU cohort (eTable 3 in the Supplement). The median TPT was 65 (IQR, 49-90) minutes for the combined cohort, with 4067 of 10 126 cases (40.2%) presenting a TPT shorter than 60 minutes. Three-quarters of patients (7434 of 10 126 [73.4%]) were transported by land, and 9098 of 10 126 (89.8%) were admitted to a designated trauma center.

In addition, TPT increased with the number of prehospital interventions performed. Indeed, a TPT lower than 60 minutes involved intubation and vasopressor use in 537 of 4049 patients (13.3%) and 228 of 4047 patients (5.6%), respectively, whereas a TPT longer than 120 minutes involved intubation and vasopressor use in 458 of 994 patients (46.1%) and 185 of 994 patients (18.6%), respectively (eTable 4 in the Supplement). Characteristics of excluded patients were similar to those who were included in the analysis, except with a slight difference in the mechanism of injury (road traffic: included patients, 5598 of 10126 [55.3%] vs excluded patients, 1551 of 3633 [42.7%]; falls: included patients, 3053 of 10126 [30.2%] vs excluded patients, 1389 of 3633 [38.2%]).

Primary Outcome

Unadjusted GLMM and GAM showed a significant, progressive increase of all-cause mortality with TPT (odds ratio [OR], 1.09 [95% CI, 1.07-1.11]; P < .001; **Figure 1**). Increasing 30-minute categories of TPT were associated with a markedly and constant increase in the risk of in-hospital death (total: 9.6% [95% CI, 9.0%-10.1%]; 0-29 minutes: 5.3% [95% CI, 3.6%-4.6%]; \geq 180 minutes, 19.4% [95% CI, 13.3%-27.3%]; **Table 2**). Odds of death increased by 9% for each 10-minute increase in TPT (OR, 1.09 [95% CI, 1.07-1.11]; **Table 3**) and by 4% after adjustment for confounders (OR, 1.04 [95% CI, 1.01-1.07]; **Figure 2**; Table 3). The post hoc power calculation demonstrated a study power of 100%.

Secondary Outcomes

Increasing 30-minute categories of TPT were associated with a progressive and constant increase in the risk of death attributable to head injury (total: 5.0% [95% CI, 4.5%-5.4%];

Figure 1. Unadjusted Models Representing the Risk of Death From All Causes, per Prehospital Time, Truncated at 200 Minutes



A, Generalized additive model: area under the curve, 0.59 (95% Cl, 0.57-0.61); internal overall calibration (expected over observed), 1.00 (95% CI, 0.94-1.06); and calibration slope, 1.03 (95% CI, 0.82-1.23). B, Generalized linear mixed model: area under the curve, 0.59 (95% CI, 0.57-0.61); internal overall

Table 2. Primary and Secondary Outcomes According to Prehospital Time



calibration (expected over observed), 1.01 (95% CI, 0.94-1.07); and calibration slope, 0.99 (95% CI, 0.79-1.19). In both graphs, the shaded area represents the 95% Cls.

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Total	Total Prehospital Total,	Total.	Overall In-Hos	Death Attributable t		
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Total Prehospital	Total,	Overall In-Hosp	IN-HOSPITAL Death Death		Death Attributable to Head Injury		Death Attributable to Bleeding	
Time, min ^a No.		Patients, No.	% (95% CI)	Patients, No.	% (95% CI)	Patients, No.	% (95% CI)	
0-29	514	27	5.3 (3.6-7.6)	13	2.5 (1.5-4.3)	8	1.6 (0.8-3.1)	
30-59	3535	244	6.9 (6.1-7.8)	129	3.7 (3.1-4.3)	53	1.5 (1.1-2.0)	
60-89	3459	344	9.9 (9.0-11.0)	179	5.2 (4.5-6.0)	63	1.8 (1.4-2.3)	
90-119	1624	208	12.8 (11.3-14.5)	109	6.7 (5.6-8.0)	36	2.2 (1.6-3.1)	
120-179	870	121	13.9 (11.8-16.4)	55	6.3 (4.9-8.1)	19	2.2 (1.4-3.4)	
≥180	124	24	19.4 (13.3-27.3)	16	12.9 (8.0-20.0)	3	2.4 (0.8-7.2)	
Total	10126	968	9.6 (9.0-10.1)	501	5.0 (4.5-5.4)	182	1.8 (1.5-2.1)	
P value for homogeneity test	NA	NA	<.001	NA	<.001	NA	.467	

Abbreviation: NA, not available.

^a Total prehospital time, arrival of the physician-led prehospital care team on scene to arrival at the hospital.

0-29 minutes: 2.5% [95% CI, 1.5%-4.3%]; ≥180 minutes, 12.9% [95% CI, 8.0%-20.0%]; Table 2). This association was also observed, although less markedly, when considering death from bleeding (total: 1.8% [95% CI, 1.5%-2.1%]; 0-29 minutes: 1.6% [95% CI, 0.8%-3.1%]; ≥180 minutes, 2.4% [95% CI, 0.8%-7.2%]; Table 2). Unadjusted odds of death attributable to head injury increased by 9% for each 10-minute increase in TPT (OR, 1.09 [95% CI, 1.07-1.11]), while unadjusted OR of death attributable to bleeding increased by 4% (1.04 [95% CI, 1.01-1.07]; Table 3). When adjusting for confounders, the association of TPT with in-hospital death attributable to head and bleeding injury showed a nonsignificant increase (Table 3).

Discussion

The main finding of this study is that longer TPT is independently associated with increasing all-cause in-hospital mor-

tality in a physician-staffed trauma system. To the best of our knowledge, this is the first study reporting such an association in a physician-staffed EMS in Europe.

The increase of mortality with increasing TPT appears to be common sense. However, existing comparative data on TPT and mortality in physician-staffed EMSs are limited and conflicting.^{20,21} Although time would be considered an independent exposure variable, it remains invariably dependent on exposure variables, such as physiology, severity of injury, and prehospital interventions. Concordantly, it was justifiable to assume that TPT could only be the result of multifactorial complexity, thereby challenging its independent association with outcome. Nevertheless, by accounting for complexity (ie, age, ISS, GCS score, and shock), time remained an independent exposure variable corroborating studies conducted in North America,²² although other studies suggest that a physician-led team exerts a positive influence on outcome,^{12,22} even when TPT is longer.

Table 3. Association Between Outcome and Prehospital Time ^a				
Death by Type	Odds Ratio by Generalized Linear Mixed Model (95% CI) ^b	P Value		
Univariable Analysis				
Overall death	1.09 (1.07-1.11)	<.001		
Death attributable to head injury	1.09 (1.06-1.11)	<.001		
Death attributable to bleeding	1.04 (1.00-1.09)	.04		
Multivariable Analysis				
Overall death	1.04 (1.01-1.07)	.002		
Death attributable to head injury	1.03 (1.00-1.07)	.15		
Death attributable to bleeding	1.00 (0.99-1.02)	.24		

^a Generalized linear model with random effect by registry and emergency medical system; adjustment for individual confounders as logarithmic function (prehospital time, age, systolic blood pressure, Injury Severity Score, and Glasgow Coma Scale score).

^bOdds ratio for increase of 10 minutes in prehospital time.

The association of TPT and mortality implies that time is an independent exposure that should be a management objective. Observed TPT were comparable with other European physician-staffed EMSs^{13,23,24} but longer than in paramedicstaffed EMS.²⁵ A propensity-matched analysis reported that critical prehospital interventions were associated with a decrease in mortality without affecting TPT.14 Clinicians constantly negotiate a trade-off between safely performing crucial interventions and limiting time on scene. The answer to this challenge may not be dichotomic concepts such as scoop and run, stay and play, or play and run. Instead, we suggest framing this challenge as a ratio between time spent on scene and interventions performed to meet the patient's critical needs. This ratio is specific to every case and prehospital situation. We suggest to call this the intervention-to-time ratio. This ratio implies that the time every prehospital intervention required must be tailored to each patient and counterbalanced by its potential gain in survival. A recent cohort study of trunk trauma⁵ revealed increased mortality in patients in shock with incompressible trunk trauma, long before any definitive hemorrhage control is possible; hence the advocacy of some people in favor of very advanced prehospital resuscitation and hemorrhage control, such as resuscitative endovascular balloon occlusion of the aorta, which carries the risk of prolonging scene time. The TPT stems from aspects crucial to prehospital care other than resuscitation, including extrication, communication with command or dispatch and the receiving center, and transport time.²⁶ The whole process needs streamlining to keep TPT as short and safe as possible, to only address the targeted and individualized essential needs of each patient that are of benefit to that patient.

When looking at the specific causes of death, the adjusted model (Table 3; Figure 2) showed a nonsignificant association of time with hemorrhage-associated mortality. The critical association of time and trauma hemorrhage was recently observed, demonstrating peak mortality from hemorrhage at 37 minutes.²⁷ This circumstance might have reduced the size of the association with death by hemorrhage in the model. Both registries only include patients who arrived alive



Multivariable generalized linear mixed model representing the risk of death from all causes according to prehospital time, adjusted for individual confounder as logarithmic function (age, systolic blood pressure, Glasgow Coma Score scale, and Injury Severity Score): area under the curve, 0.96 (95% CI, 0.95-0.96); internal overall calibration (expected over observed), 1.00 (95% CI, 0.96-1.04); and calibration slope, 1.00 (95% CI, 0.94-1.06). The full model is presented in eTable 5 in the Supplement. The shaded area represents the 95% CIs.

at the hospital. Ball et al²⁸ indicated that with shorter TPT, more patients arrived alive but then died subsequently. Future studies need to address this question and include all patients deceased on scene.

In opposition to previous reports,²⁹ the adjusted model did not show a significant association of TPT and TBI-induced mortality. Affirmation of TBI as the cause of death is difficult since patients with head injuries commonly have other confounding causes of death (eg, hemorrhage, thoracic trauma). Moreover, a substantial proportion of patients with TBI ultimately die after withdrawal of treatment and are registered as such, which may have biased the true proportion of death attributable to TBI.

This study has several strengths. Data were extracted from 2 established registries with robust data management. These data were collected from 2 geographically and structurally distinct regions with different trauma systems. The TRENAU system qualifies as inclusive, with 3 designated levels of care.¹⁶ The Paris system tends to be exclusive, with 6 designated level I centers absorbing the most regional trauma cases.¹⁷ These characteristics apply to other European EMS, supporting the external validity and generalizability of the observations made. Thus, the present model appears robust and generalizable. As eFigure 2 in the Supplement illustrates, mortality retains a coherent and consistent association to well-documented variables and confounders.

Limitations

This study carries a number of limitations linked to its design. Based on a prospective data collection and with adequate control of potential confounders, it appeared appropriate, however, to examine the hypothesis in a sufficiently large real-world cohort. A physician closing the case determined the cause of death, and a subjective bias cannot be excluded, which is in analogy to most existing registries, unless postmortem computed tomographic scans and/or systematic autopsies become standard. Only patients who arrived alive in the hospital were included in the study, a feature shared with most established trauma registries that may have reduced the sign of the association of hemorrhage-induced mortality and time. The choice of adjustment variables may seem arbitrary and the number of variables limited. However, the adjustment aimed to compensate for case complexity and case mix in a simple but robust way. Patients with shock (as indicated by systolic blood pressure <90 mm Hg) and/or TBI (low GCS scores) and high injury load (high ISS) and elderly patients account for the greater part of this complexity, since they may require prehospital resuscitation. Hence, the models did not include any specific prehospital interventions, because they are well correlated and share a high collinearity with any of the 4 adjustment variables in the model and are part of the socalled causal pathway.

Prehospital interventions in France focus on advanced airway management, fluid and vasoactive therapy, and simple hemostatic measures (eTable 4 in the Supplement). At the time of the study, blood product use was rare, and advanced surgical techniques were exceptional.

We decided to perform complete case analysis instead of performing multiple imputation. We cannot exclude that exclusion of observations led to selection bias. Multiple imputation is not always superior to complete case analysis; in particular, if missing data are not missing at random, imputation may generate a bias toward the null.³⁰ In this study, it seems inappropriate to impute data on prehospital time without precise geographical location, because ambulances in France are currently not systematically equipped with geographic positioning systems.

Missing data led to the exclusion of 23% of patients from both registries. However, apart from a slight difference in the mechanism of injury, patients who were excluded shared the characteristics and severity of injury similar to the patients who were included. This study included cases from a large urban and a rural, mountainous area, supporting the generalizability of the results to most of European and mature trauma systems. A cluster effect cannot be excluded. The use of a mixed model with a random effect should however control for this bias. These results may less apply to very remote isolated regions and or systems with a higher incidence of penetrating injury.

Conclusions

This study demonstrates an independent association of in-hospital, all-cause mortality in trauma with total prehospital time in a physician-staffed EMS in France, even after adjustment for case complexity. These results imply that prehospital time management should become a management objective. Both findings plead for further streamlining of prehospital trauma care and the need to define the optimal intervention-to-time ratio.

ARTICLE INFORMATION

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Supplementary Online Content

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This supplementary material has been provided by the authors to give readers additional information about their work.

eTable 1. Complete list of collected variables

Demographics	Prehospital	Hospital	Outcome and follow-
			ир
Age	MICU team	Hospital name	In-hospital death
Gender	Glasgow coma scale	Hospital level	Cause of death
Mechanism of injury	Blood pressure	Glasgow coma scale	Length of stay (ICU)
Circumstances	Respiratory rate	Blood pressure	Length of ventilation
Place of accident	Heart rate	Respiratory rate	Withdrawal of treatment
Date of accident	Pulse oximetry	Heart rate	Abbreviated injury scale
	Pupil size	Pulse oximetry	Injury severity score
	Intubation	Pupil size	
	Fluid resuscitation	Intubation	
	Vasopressor use	Fluid resuscitation	
	Osmotherapy	Vasopressor use	
	Arrival on scene (time)	Osmotherapy	
	Arrival at hospital (time)	Surgery	
	Transportation type	Treatment	
	Cardiac arrest	Para-clinical tests	

	Considered for analysis	Not considered for analysis due to missing data on
	N=10,126	prehospital time
		N=3,633
Age, mean (95%CI)	41 (40-41)	41 (40-41)
Sex male, N; % (95%CI)	7,937; 78% (78-79)	2,828; 78% (76-79)
Penetrating injury, N; % (95%CI)	861; 9% (8-9)	351; 10% (9-11)
Road traffic injury, N; % (95%CI)	5,598; 55%(54-56)	1,551; 43% (41-44)
Fall, N; % (95%Cl)	3,053; 30% (29-31)	1,389; 38% (37-40)
SBP mean (95%CI)	117 (116-117)	120 (119-121)
GCS, mean (95%CI)	13 (13-13)	13 (13-13)
N ;% (95%CI)		
3-8	1,518; 15% (14-16)	497; 15% (14-17)
9-13	963; 10% (9-10)	293; 9% (8-10)
14-15	7,453; 74% (73-74)	2,461; 76% (74-77)
Prehospital intubation N ;%	2,321; 23% (22-24)	704; 19% (18-21)
(95%CI)		
Haemorrhagic shock, N ;%	1,010; 10% (9-11)	287; 8% (7-9)
(95%CI)		
AIS Head ≥3, N ;% (95%CI)	2,724; 27% (26-28)	1,053; 29% (28-30)
Mean ISS	17 (17-18)	17 (16-17)
ISS >15, N ;% (95%CI)	4,757; 47% (46-48)	1,692; 47% (45-48)
In-hospital Death, N ;% (95%CI)	968; 10% (9-10)	291; 8% (7-9)

eTable 2. Characteristics of patients excluded for analysis and missing data

AIS=abbreviated injury scale, GCS=Glasgow Coma Scale, ISS=Injury severity score, SBP=systolic blood pressure

	Total	Paris_ lle de	Northern French
		France	Alps (TRENAU)
		(TRAUMABASE)	
	N=10,126	N=5,067	N=5,059
Ambulance transportation, N (%)	7,434 (76)	4,319 (89)	3,115 (63)
Helicopter transportation, N (%)	2,367 (24)	516 (11)	1,851 (37)
Transport to designated trauma	9,098 (90)	5,067 (100)	4,031 (80)
center, N (%)			
Median TPT [IQR]	65 [49-90]	73 [54-100]	60 [45-80]
Total Prehospital time ^a , N (%)			
0-29 min	514 (5)	247 (5)	267 (5)
30-59 min	3,535 (35)	1,335 (26)	2,200 (43)
60-89 min	3,459 (34)	1,762 (35)	1,697 (34)
90-119 min	1,624 (16)	999 (20)	625 (12)
120-179min	870 (9)	630 (12)	240 (5)
180 + min	124 (1)	94 (2)	30 (1)

eTable 3. Transportation mode and time according to regional database.

^aTPT = Total Prehospital Time, arrival of the physician-lead prehospital care-team on scene to arrival at the hospital

	Total	Oro-tracheal	Resuscitation	Vasopressor	Osmotherapy
		intubation	fluid >1000 ml	use	N (%)
		N (%)	N (%)		
				N (%)	
0-29 min	514	49 (10)	71 (14)	24 (5)	11 (2)
30-59 min	3,535	488 (14)	578 (16)	203 (6)	131 (4)
60-89 min	3,459	794 (23)	806 (23)	328 (10)	210 (6)
90-119 min	1,624	555 (34)	550 (34)	242 (15)	123 (8)
120-	870	391 (45)	365 (42)	153 (18)	76 (9)
179min					
180 + min	124	67 (54)	68 (54)	32 (26)	18 (14)
Total	10,126	2,321 (23)	2,438 (24)	982 (10)	569 (6)
P value for		<0.001	<0.001	<0.001	<0.001
Homogenei					
ty test					

eTable 4. Pre-hospital Interventions and Total Pre-hospital Time

^aTPT = Total Prehospital Time, arrival of the physician-lead prehospital care-team on scene to arrival at the hospital

eTable 5. Generalized Linear Mixed Model with primary outcome all-causes of death.

	Estimate	Standard	P-Value	Odds ratio	95% CI
		error			
Prehospital time	0.0389	0.0145	0.007	1.040	1.011 – 1.069
(by 10 minutes)					
Age	-0.0366	0.0131	0.005	0.964	0.940-0.989
Age^2	0.0090	0.0001	<0.001	1.001	1.001-1.001
SBP	-0.0399	0.0074	<0.001	0.961	0.947-0.975
SBP^2	3.6 e ⁻⁵	8.9e ⁻⁵	0.687	1.00	0.999-1.000
SBP^3	6.9 e ⁻⁷	2.9e ⁻⁷	0.02	1.00	1.00-1.00
ISS	0.1789	0.0342	<0.001	1.196	1.130 – 1.265
ISS^2	-0.0032	8.6 e ⁻⁴	<0.001	0.997	0.995-0.998
ISS^3	2.5 e ⁻⁵	7.6 e ⁻⁶	0.001	1.00	1.00 -1.00
GCS	-0.0260	0.0083	<0.001	0.786	0.770 – 0.803

^2: quadratic term; ^3 cubic term; SBP: Systolic blood pressure; ISS: Injury Severity Score; GCS: Glasgow Coma Scale. Random effect by Registry and EMS system

eTable 6. Generalized Linear Mixed Model with primary outcome all-causes of death.

	OR (95% CI)	P value
TPT by 10 min	1.04 (1.01-1.07)	0.002
Age	0.96 (0.95-0.99)	0.006
Age ²	1.001 (1.00-1.001)	<0.001
SBP	0.95 (0.94-0.96)	<0.001
SBP ²	1.00 (1.00-1.00)	<0.001
GCS	0.56 (0.49-0.65)	<0.001
GCS ²	1.02 (1.01-1.02)	<0.001
ISS	1.18 (1.11-1.25)	<0.001
ISS ²	0.99 (0.99-0.99)	0.002
ISS ³	1.00 (1.00-1.00)	0.006

TPT : Total Prehospital Time, SBP :Systolic Blood Pressure ; GCS ; Glasgow Coma Scale ; ISS :Injury Severity Scale.²indicates quadratic term ³indicates cubic terms.

eFigure 1. Flowchart of study

Fig 1. Study flow chart based on the Northern French Alps and Paris-Ile-de-France Registries.



MICU: Mobile Intensive Care Unit

*Mountain rescue: Speleology (1), Pedestrian transportation (2), Meteorological concerns (2)



eFigure 2. Adjusted Generalized Additive Model for individual confounders as logarithmic function (prehospital time, age, systolic blood pressure, Injury Severity Score, GCS); the Y-axis represents the functional risk of death